

# Spatial sign tests for canonical correlations and independence between two vectors

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## 1 Introduction

Let

$$\mathbf{z} = (\mathbf{z}_1, \dots, \mathbf{z}_n)^T = \begin{pmatrix} \mathbf{x}_1 & \dots & \mathbf{x}_n \\ \mathbf{y}_1 & \dots & \mathbf{y}_n \end{pmatrix}^T = \begin{pmatrix} \mathbf{x} \\ \mathbf{y} \end{pmatrix}^T$$

be a  $k$ -variate data set of  $n$  observations from distribution with cdf  $F$  and partitioned covariance matrix

$$\Sigma = \Sigma(F) = \begin{pmatrix} \Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & \Sigma_{22} \end{pmatrix}.$$

We wish to test the null hypothesis  $H_0^1: \Sigma_{12} = 0$  or the null hypothesis  $H_0^2: \mathbf{x}$  and  $\mathbf{y}$  are independent. Note that  $H_0^1$  implies that all the canonical correlations are zero, but does not imply the independence between two vectors unless their joint distribution is multinormal.

Several tests for correlations and independence are found in the literature. Wilks (1935) derived the likelihood ratio test for  $H_0: \Sigma_{12} = 0$  when the observations are multinormal. An other classical test statistic is introduced for example by Pillai (1955). If  $S = S(\mathbf{z})$  is a partitioned sample covariance matrix, then the Pillai's test statistic is expressed as  $T_0 = \text{Tr}(S_{11}^{-1}S_{12}S_{22}^{-1}S_{21})$ . Nonparametric approaches to the problem were given by Puri and Sen (1971) and Gieser and Randles (1997). The latter based their test on interdirections. They found limiting distribution of the test statistic under the null hypothesis of independence and under contiguous alternatives and compared the efficiency and robustness of their test with the classical ones.

## 2 Tests based on signs

Assume now that  $F$  is elliptic. For testing  $H_0^1$ , a shape matrix  $V = V(\mathbf{z})$  based on standardized spatial signs (Randles, 2000, Hettmansperger and Randles, 2001) is first constructed: The spatial sign of  $\mathbf{z}$  is  $\mathbf{S}(\mathbf{z}) = \|\mathbf{z}\|^{-1}\mathbf{z}$ , that is, the unit vector in the direction of  $\mathbf{z}$ . The location and shape estimates  $\hat{\boldsymbol{\mu}}$  and  $\hat{V}$  are chosen so that the spatial signs of the transformed observations  $\hat{V}^{-1/2}(\mathbf{z}_i - \hat{\boldsymbol{\mu}})$ ,  $i = 1, \dots, n$  are standardized, that is,

$$\text{ave}_i\{\hat{\mathbf{S}}_i\} = \mathbf{0} \quad \text{and} \quad \text{ave}_i\{\hat{\mathbf{S}}_i\hat{\mathbf{S}}_i^T\} = \frac{1}{k}I_k.$$

Note that  $\hat{V}$  is known as Tyler's M-estimate (1987). The test statistic is then  $T_1 = \text{Tr}(V_{11}^{-1}V_{12}V_{22}^{-1}V_{21})$  and the standardized version of the test statistic has a limiting chi-squared distribution.

For testing  $H_0^2$ , standardized spatial sign vectors  $S_i$ ,  $i = 1, \dots, n$  for  $\mathbf{x}$  and  $R_i$ ,  $i = 1, \dots, n$  for  $\mathbf{y}$  are constructed separately. Then  $H_{12} = \text{ave}_i\{S_i R_i^T\}$  is the covariance matrix between  $\mathbf{x}$ - and  $\mathbf{y}$ -signs and the test statistic is  $T_2 = \text{Tr}(H_{12}H_{12}^T)$ . Again (under general assumptions) the limiting distribution of the test statistic is chi-squared distribution.

The tests are compared with the classical ones through limiting Pitman efficiencies in the multivariate normal and  $t$ -distribution cases. The theory is illustrated by an example.

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